

# Mean Field Behavior during the Big Bang for Coalescing Random Walk

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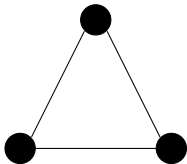
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Motivation: duality with the voter model.

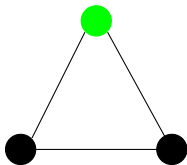
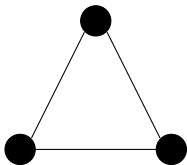
## An example

Black=occupied, Green =vacant



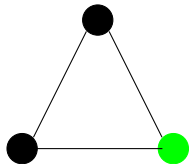
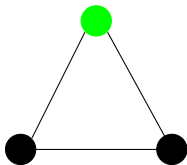
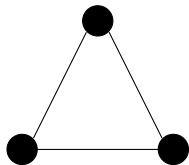
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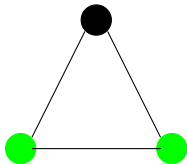
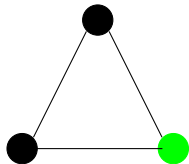
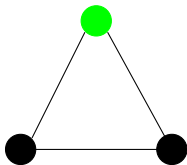
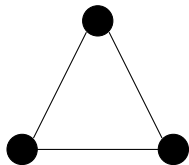
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## CRW on the complete graph

$G$  is a complete graph (clique).  $r_{x,y} = 1/(n-1)$ .

$L_t$ : # of walkers at time  $t$ .

$L_0 = n$ .  $L_t \rightarrow L_t - 1$  at rate  $L_t(L_t - 1)/(n-1)$ .

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$$\tau_{\text{coal}} = \sum_{i=1}^n \frac{e_i}{i(i-1)/n}.$$

- $e_i, i \geq 1$  are i.i.d. with dist.  $\text{Exp}(1)$ .
- $\frac{e_i}{i(i-1)/n}$  is the time it takes for the  $n-i+1$ -th coalescence to occur (corresponding to  $L_t$  from  $i$  to  $i-1$ ).

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Related model: Kingman's coalescent.  $L_0 = \infty$ .  $L_t \rightarrow L_t - 1$  at rate  $L_t(L_t - 1)/2$ .

## Decay of density on the complete graph

Define the expected density (expected fraction of occupied sites)

$$P_t = \frac{\mathbb{E}(L_t)}{n}.$$

Determine  $L_t$ : the time it takes to make  $h$  coalescences

$$\sum_{i=n-h+1}^n \frac{e_i}{i(i-1)/(n-1)} \sim n \left( \frac{1}{n-h} - \frac{1}{n} \right)$$

for  $1 \ll h \ll n$ . Set this expression to be  $t$ , we get

$$L_t = n - h \sim \frac{n}{t+1}.$$

Thus

$$P_t \sim \frac{1}{t+1}.$$

# Spatial structure

Often there is a spatial structure.

- $\mathbb{Z}^d$ .
- $\mathbb{T}^d$ .
- General vertex transitive graphs.
- Random graphs (e.g., configuration model).

## Heuristic argument [van den Berg-Kesten, 2000]

Consider  $\mathbb{Z}^d$ .  $P_t = P_t(o)$ : prob. that origin is occupied at time  $t$ .  
Take  $1 \ll \Delta(t) \ll t$ .

$$\begin{aligned} -\frac{dP_t}{dt} &= \mathbb{P}(o \text{ and } \mathbf{e}_1 \text{ occupied at } t) \\ &\sim \sum_{x,y} \mathbb{P}(x \text{ and } y \text{ occupied at } t - \Delta(t)) \times \\ &\quad \mathbb{P}(x + S_{\Delta(t)} = o, y + S'_{\Delta(t)} = \mathbf{e}_1, x + S_r \neq y + S'_r, \forall r \leq \Delta(t)) \\ &\sim P_{t-\Delta(t)}^2 \alpha_{\Delta(t)}. \end{aligned}$$

- $x$  and  $y$  are the location of the walkers that later come to  $o$  and  $\mathbf{e}_1$ .  $S, S'$ : independent random walks starting from  $o$ .
- $\alpha_{\Delta(t)}$ : the probability that two time-reversed random walk starting from  $o$  and  $\mathbf{e}_1$  don't collide by time  $\Delta(t)$ .

## Results on $\mathbb{Z}^d$

Assuming  $P_t \sim P_{t-\Delta(t)}$  and  $\alpha_t \sim \alpha_{t-\Delta(t)}$ . The heuristic suggests that  $P_t \approx 1/(t\alpha_t)$  for moderately large  $t$ . This was known to be true for SRW on  $\mathbb{Z}^d$ ,  $d \geq 2$ .

### Theorem (Bramson-Griffeath, 1980)

Consider the CRW on  $\mathbb{Z}^d$ . We have, as  $t \rightarrow \infty$ ,

$$P_t \sim \begin{cases} \frac{1}{\pi} \frac{\log t}{t} & d = 2 \\ (\gamma_d t)^{-1} & d \geq 3 \end{cases}$$

where  $\gamma_d$  is the probability that a simple random walk in  $\mathbb{Z}^d$  starting from origin never returns to it.



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By justifying previous heuristic argument, [van der Berg-Kesten, 2000] proved the same result for  $d \geq 3$ .

## Approximation for coalescence time

$\pi$ : stationary distribution.

Mean meeting time (the time it take for two indep. walkers to meet)

$$t_{\text{meet}} = \mathbb{E}_{\pi^2}(\tau_{\text{meet}}).$$

For complete graph  $t_{\text{meet}} = (n - 1)/2$ .

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Aldous and Fill conjectured that for finite transitive graph (transitivity means the graph looks the same from every vertex)

$$\frac{\tau_{\text{coal}}}{t_{\text{meet}}} \sim \sum_{i=2}^{\infty} \frac{e_i}{i(i-1)/2}.$$

Equality holds for complete graph (replacing  $\infty$  by  $n$ ).  $e_i \sim \text{Exp}(1)$ .

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The factor  $i(i-1)/2$  counts the number of pairs

- Why exponential?

## Aldous-Brown approximation

### Lemma (Aldous-Brown, 1992)

*For an irreducible reversible Markov chain on a finite state  $V$  with stationary distribution  $\pi$  and  $A \subset V$ , if we denote the hitting time of  $A$  by  $T_A$  and its density function w.r.t. the stationary chain by  $f_{T_A}$ , then*

$$\left| \mathbb{P}_\pi(T_A > t) - \exp\left(-\frac{t}{\mathbb{E}_\pi(T_A)}\right) \right| \leq \frac{t_{\text{rel}}}{\mathbb{E}_\pi(T_A)},$$

and

$$\frac{1}{\mathbb{E}_\pi(T_A)} \left(1 - \frac{2t_{\text{rel}} + t}{\mathbb{E}_\pi(T_A)}\right) \leq f_{T_A}(t) \leq \frac{1}{\mathbb{E}_\pi(T_A)} \left(1 + \frac{t_{\text{rel}}}{2t}\right).$$

Consider the product chain and take  $A$  to be the diagonal set. We have  $\mathbb{E}_\pi(T_A) = t_{\text{meet}}$ .

## Second Prediction

[Oliveira, 2013] proved the Aldous-Fill conjecture under the condition  $t_{\text{mix}} \ll t_{\text{meet}}$  (equivalent to  $t_{\text{rel}} \ll t_{\text{meet}}$  due to Hermon).  $t_{\text{mix}}$  and  $t_{\text{rel}}$  quantify the rate of convergence to stationary distribution (See Markov Chains and Mixing Times).

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$$t_{\text{meet}} \sum_{i \geq n-h+1} \frac{e_i}{i(i-1)/2} \sim \frac{2t_{\text{meet}}}{n-h}.$$

$$\frac{2t_{\text{meet}}}{n-h} = t \Rightarrow n-h = \frac{2t_{\text{meet}}}{t}.$$

Hence we have another prediction

$$P_t = \frac{E(L_t)}{n} = \frac{n-h}{n} \sim \frac{2t_{\text{meet}}}{nt}.$$

## Equivalence of the two predictions

Two predictions for  $P_t$



$$P_t \sim \frac{1}{t\alpha_t}$$

where  $\alpha_t = r(o)\mathbb{P}_{o,\nu_o}(\tau_{\text{meet}} > t)$  ( $\nu_o$  is a random neighbor of  $o$ )



$$P_t \sim \frac{2t_{\text{meet}}}{nt} \text{ for finite graphs}$$

They are equivalent to each other for many graphs by Kac's formula (in continuous time) and Aldous-Brown approximation:

$$\frac{1}{t_{\text{meet}}} \sim f_{T_A}(t) = \frac{2}{n}\mathbb{P}_{o,\nu_o}(\tau_{\text{meet}} > t) \text{ for } r(o) = 1.$$



## Main Results: finite graphs

### Theorem (Hermon-Li-Yao-Zhang, 2021)

*Two predictions holds as long as  $1 \ll t \ll t_{coal}$  (called the Big Bang regime since the number of particles is evolving rapidly in this regime) for*

- *transitive graphs  $G_n$  such that  $\text{diam}(G_n)^2 \ll n/\log n$ ,*
- *Configuration Model  $\mathbb{CM}(n, D)$  with  $3 \leq D < M$ .  
If  $D$  is a constant  $d$  then  $\mathbb{CM}(n, D)$  is random  $d$ -regular graph.*

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Remarks:

- For such graphs  $t_{\text{coal}}$  and  $t_{\text{meet}}$  both have order  $n$ .
- By [Tessera and Tointon, 2019],  $\text{diam}(G_n)^2 \ll n/\log n$  implies

$$\lim_{s \rightarrow \infty} \limsup_{n \rightarrow \infty} \sup_{x, y} \int_{s \wedge t_{\text{rel}}}^{t_{\text{rel}}} p_s(x, y) ds = 0.$$

## Configuration model

Construction of the configuration model  $\mathbb{CM}_n(D)$

- Let  $D$  be a probability measure on  $\mathbb{Z}_+$ , and  $n \in \mathbb{Z}_+$ .
- We take  $n$  vertices labeled  $1, \dots, n$ , and  $d_1, \dots, d_n$  i.i.d. sampled from  $D$ .
- For each vertex  $i$  we attach  $d_i$  half edges to it. Then we get  $G_n$  by uniformly matching all half edges, conditioned on  $\sum_{i=1}^n d_i$  being even.

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The *local weak limit*  $\text{UGT}(D)$  of  $\mathbb{CM}_n(D)$  is a unimodular Galton-Watson tree where

- the root has offspring distribution  $D$
- later generations have offspring distribution  $D^*$ :

$$\mathbb{P}(D^* = k) := \frac{(k+1)\mathbb{P}(D = k+1)}{\sum_{i=0}^{\infty} i\mathbb{P}(D = i)}$$

# Main Results: infinite Graphs

Theorem (Hermon-Li-Yao-Zhang, 2021)

The prediction  $P_t(o) \sim 1/(t^\alpha)$  as  $t \rightarrow \infty$  where

$$\alpha = \mathbb{E}(r(o)\mathbb{P}_{o,\nu_o}(\tau_{meet} = \infty))$$

holds for

- all transient transitive unimodular graphs, including
  - Cayley graphs
  - amenable graphs(=graphs with subexponential decay of return probability)
- unimodular Galton-Watson tree  $UGT(D)$ . If  $D$  is a constant  $d$  then  $UGT(D) = \mathbb{T}^d$ .

## Duality with voter model

Voter model: at rate  $r_{x,y}$ ,  $x$  adopts the opinion of  $y$ .

A site is occupied in CRW  $\leftrightarrow$  the opinion is not lost in VM.

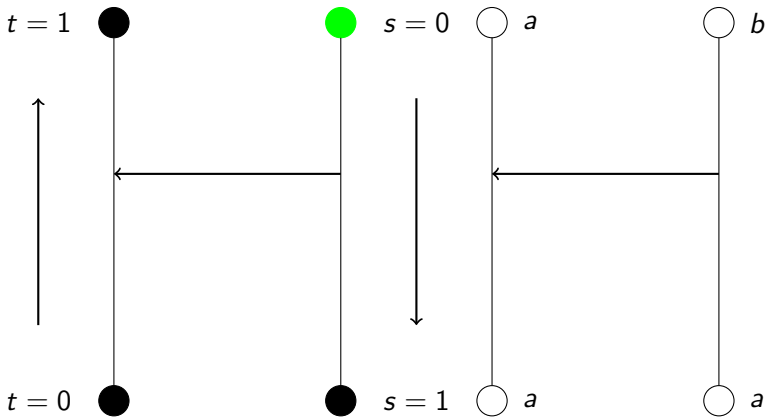


Figure: Left panel: CRW; right panel: voter model

## Proof Sketch of [Bramson-Griffeath,1980]

$n_t$ : # walkers that end up at origin at time  $t$ .

$\eta_t$ : the voter model starting from different opinions at every site.

$\hat{N}_t := \{x : \eta_t(x) = \eta_t(o)\}$ . [Kelly, 1977] gives

$$\mathbb{P}(\hat{N}_t = j) = j\mathbb{P}(n_t = j), j \geq 0, \text{ ( i.e., size-biased version of } n_t)$$

$$P_t = \mathbb{P}(n_t > 0) = \mathbb{E}(\hat{N}_t^{-1}) = \mathbb{E} \left[ \left( \frac{\hat{N}_t}{\mathbb{E}(\hat{N}_t)} \right)^{-1} \right] \mathbb{E}(\hat{N}_t)^{-1}.$$

$\mathbb{E}(\hat{N}_t)$  is equal to  $\mathbb{E}(R_{2t})$  where  $R$  is the range of a random walk.

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### Theorem (Sawyer, 1979)

Consider CRW on  $\mathbb{Z}^d, d \geq 2$ .

$$\lim_{t \rightarrow \infty} \mathbb{E} \left[ \left( \frac{\hat{N}_t}{\mathbb{E}(\hat{N}_t)} \right)^k \right] = \frac{(k+1)!}{2^k}.$$



## Proof Sketch of [Bramson-Griffeath,1980]-cont'd

A remark from [Bramson-Griffeath,1980]: Sawyer's theorem comes tantalizingly close to determining the asymptotics of  $P_t$ .

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Theorem (Bramson-Griffeath, 1980)

$$P_t = \begin{cases} O\left(\frac{\log t}{t}\right) & d = 2, \\ O\left(\frac{1}{t}\right) & d \geq 3. \end{cases}$$

Lemma (Bramson-Griffeath, 1980)

*Sawyer's Theorem+upper bound on  $P_t$  gives the asymptotics of  $P_t$ .*

Basically, the upper bound on  $P_t$  implies the  $\hat{N}_t/E(\hat{N}_t)$  doesn't have too much mass near 0.

## Transform to coalescence probability

Let  $N_t$  be the number of walkers that collide with the walker starting at  $U$ .  $N_0 = 1$ .

$$N_t = \sum_x \mathbf{1}[\text{the particle starting at } x \text{ coalesced with } U \text{ by time } t]$$

$$P_t = \mathbb{E}(N_t^{-1}) = [\mathbb{E}(N_t)]^{-1} \mathbb{E} \left[ \left( \frac{N_t}{\mathbb{E}(N_t)} \right)^{-1} \right].$$

(A graph rooted at a uniform vertex is unimodular.)

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(A graph rooted at a uniform vertex is unimodular.)

$\mathbb{C}$ : coalescence time for  $k+1$  walkers.

$$\begin{aligned} \mathbb{E}(N_t^k) &= \frac{1}{n} \sum_{x_1, \dots, x_{k+1} \in V} \mathbb{E}(\mathbf{1}[X_i(0) = x_i, \forall 1 \leq i \leq k+1]) \\ &\quad \times \mathbf{1}[\mathbb{C}(X_1, \dots, X_{k+1}) \leq t] \\ &= n^k \mathbb{P}_{\pi \otimes k+1}(\mathbb{C}(X_1, \dots, X_{k+1}) \leq t), \end{aligned}$$

## Ingredients of the proof

Using the machinery in the proof of  $\mathbb{Z}^d$  case by Braomson-Griffeath, it suffices to

- give an upper bound of  $P_t$  that differs from the 'true value' of  $P_t$  by a multiplicative constant,
- show that the coalescence probability

$$\mathbb{P}_{\pi^{k+1}}(\mathbb{C}(X_1, \dots, X_{k+1}) \leq t) \sim (k+1)! \left( \frac{t}{t_{\text{meet}}} \right)^k.$$

Another indication of mean field! B-G proof heavily relies on the specific geometric structure of  $\mathbb{Z}^d$ .

## Solution

- For the first part, we show that for any  $t > 0$ ,

$$c \frac{\inf_{x \in G} \int_0^t p_s(x, x) ds}{t} \leq P_t \leq C \frac{\sup_{x \in G} \int_0^t p_s(x, x) ds}{t}.$$

where  $c$  and  $C$  are universal constants.

- For the second part, we use the reversibility of random walk to transform collision probability to non-colliding probability. If two forward paths collide at  $t$  then (after reversing time) the backward paths don't collide in  $[0, t]$ .

We want to estimate  $\mathbb{P}_{\pi \otimes k+1}(\mathbb{C}(X_1, \dots, X_{k+1}) \leq t)$ .

Consider the case  $k = 1$ . The probability that  $X_1$  and  $X_2$  collide within time interval  $[t, t + dt]$  is about

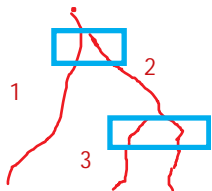
$$\begin{aligned} & 2 \sum_{u,v} \mathbb{P}(X_1(t) = u, X_2 \text{ jumps from } u \text{ to } v \text{ in } [t, t + dt]) \\ & \sim 2 \sum_{u,v} \mathbb{P}(X_1(t) = u, X_2(t) = v, \text{ no collisions in } [0, t]) r_{v,u} dt \\ & \sim 2 \sum_u \mathbb{P}(\gamma_1(0) = u) r(u) \times \\ & \quad \sum_v \frac{r_{u,v}}{r(u)} P(\gamma_2(0) = v) \mathbb{P}_{u,v}(\gamma_1(s) \neq \gamma_2(s), \forall 0 \leq s \leq t) dt, \end{aligned}$$

where  $\gamma_1$  and  $\gamma_2$  are the time-reversals of  $X_1, X_2$  on  $[0, t]$ .

# Collision Pattern and Branching Structure

We can imagine  $\gamma_1$  is the parent of  $\gamma_2$  and interpret the term  $r_{u,v}/r(u)$  as the probability of the particle at  $u$  giving birth to a particle at location  $v$ .

Can be generalized to  $k \geq 3$  paths.





If two walkers don't collide in time  $O(t_{\text{rel}})$ , then they will also not collide before time  $t$ .

### Lemma

*For any  $x \neq y$  and  $0 < s < t$ , the probability that two walkers starting from  $x$  and  $y$  collide between time  $s$  and  $t$  is bounded by*

$$2 \exp(-s/t_{\text{rel}}) \frac{\max_z \int_0^{2s} p_s(z, z) ds}{\min_z \int_0^{2s} p_s(z, z) ds} + \frac{8t}{n} (s^{-1} \vee r_{\text{max}})$$

$r_{\text{max}} = \max_x r(x)$ . The error is small for  $t_{\text{rel}} \ll s \leq t \ll n$ .

## Open Question

For finite graphs our results (the expectation of the number of occupied sites) can be upgraded to a weak law of large numbers using negative correlation

$$\mathbb{P}(\text{both } x \text{ and } y \text{ occupied at } t) \leq \mathbb{P}(x \text{ occupied at } t)\mathbb{P}(y \text{ occupied at } t).$$

What about fluctuations? Do we have a Gaussian limit as in the mean field case ([Aldous, 1999])?

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Thanks!